

Lecture 6

Introduction to queuing systems

6.1 Kendall's classification

D.G. Kendall (1951) has introduced the following classification of the problems at hand. Three symbols are given, separated by two oblique strokes: $-/-/$. The last symbol specifies the number of servers. The first and the second denote the type of distribution of inter-arrival time and of the holding time (service time), respectively. An exponential distribution is indicated by the symbol M (called after the "Markovian property", which essentially is the property of forgetfulness). Constant intervals are characterized by the symbol D (deterministic). Arbitrary distribution are given the symbol G ('general'). If the later case refers to inter-arrival times and if, moreover, successive inter-arrivals are supposed to be independent, this is indicated by the symbol GI (general independent). Some examples:

- $M/M/c$ means: Poisson arrival process, exponential distribution of holding time, c servers,
- $M/G/1$ (sometimes we write also $M/GI/1$ where GI stands for general independent) means Poisson arrival process, arbitrary distribution of holding times one server,
- $GI/D/1$ means inter-arrival times 'general independent', constant holding time, one server.

In the cases to be dealt with, the first symbol will nearly always be M (Poisson). Blocking and delay systems will be distinguished by simply writing 'blocking' or 'delay' behind classification.

6.2 Standard systems

By “standard cases“ it is meant the simple steady-state cases M/M/c-blocking and M/M/c-delay. The arrival process is Poissonian in both cases and the distribution of holding-times is exponential. Moreover, it is assumed that the order of serving delayed demands is not dependent on a prior knowledge of holding-times.

6.2.1 Probability of blocking for the system M/M/c-blocking

The arrival process is Poissonian with arrival (or birth) rate ρ . There are c servers. The holding-time is exponential (1). There is no waiting facility. If all servers are occupied, arriving demands are discarded.

When the number of busy servers is r ($r = 0, 1, \dots, c$) the system will be said to be “in state $[r]$ “. It is assumed that stationarity exists. This means that the probability of the system being in $[r]$ is time-independent:

$$p_r := Pr\{[r] \text{ at time } \tau\} = \text{const.}$$

At an arrival (rate= ρ) the index r of the state of the system is increased by as long as $r < c$. If $r = c$ an arrival is ineffective. When the system is in $[r]$ the arrival rate remains ρ , as for the Poisson process the inferential knowledge conveyed to us by the specification of the state is irrelevant. The total termination rate is r ; possible inferential knowledge about the ages of the r occupations is irrelevant as the holding-time is exponential. The probability of the state and rates of arrival and termination (collectively called transitions) may be summarized in the following scheme:

$$[0] \xrightleftharpoons[\rho]{\rho} [1] \xrightleftharpoons[\rho]{\rho} \dots [r-1] \xrightleftharpoons[\rho]{\rho} [r] \xrightleftharpoons[\rho]{\rho} [r+1] \dots [c-1] \xrightleftharpoons[\rho]{\rho} [c],$$

probability of the states are respectively $p_0, p_1, \dots, p_{r-1}, p_r, \dots, p_{c-1}, p_c$.

Now, stationarity requires that the probability of the formation of $[r]$ in $\Delta\tau$ should equal the probability of a departure from $[r]$ during $\Delta\tau$. The formation is possible from state $[r-1]$, which has probability p_{r-1} , by an arrival, which has probability $\rho\Delta\tau$. The joint probability of these events is $p_{r-1}\rho\Delta\tau$. When $r < c$, another possibility is from probability $[r+1]$ by "a death of one out of $r+1$ " (conditional probability $(r+1)\Delta\tau$): joint probability $p_{r+1}(r+1)\Delta\tau$. Formations by multiple transitions have probabilities of order $O(\Delta\tau^2)$ and are discarded. When the system is in $[r]$, where $r < c$, both arrival and termination would mean departure from $[r]$; hence the probability of a departure from $[r]$ in $\Delta\tau$ is: $p_r(\rho+r)\Delta\tau$. So

$$p_{r-1}\rho\Delta\tau + p_{r+1}(r+1)\Delta\tau = p_r(\rho+r)\Delta\tau.$$

or

$$\rho p_{r-1} - (\rho+r)p_r + (r+1)p_{r+1} = 0,$$

$r = 0, 1, \dots, c - 1$.

For this equation to hold for $r = 0$ too, p_{-1} should be taken equal to 0. For the state $[c]$ it should be observed that no transitions to and from a state $[c + 1]$ are possible. Hence,

$$\rho p_{c-1} - cp_c = 0.$$

The above mentioned set of equations are called *birth-and death equations*. It is a set of $c + 1$ homogeneous linear equations in $c + 1$ unknowns. So, for a nontrivial solution to exist, the equations must be linearly independent. By adding these equations for $0, 1, \dots, r$ one obtains

$$-\rho p_r + (r + 1)p_{r+1} = 0. \quad (6.1)$$

the last of which is identical to previous equation, showing the dependency.

The set of equations. (6.1) should be supplemented by the statement that the defined states form a complete set of mutually exclusive events, and hence,

$$\sum_{r=0}^c p_r = 1 \quad (6.2)$$

The equations. (6.1) taken consecutively for $r = 0, 1, 2, \dots$ enable one to express p_1 in terms of p_0 , p_2 in terms of p_1, \dots, p_c in terms of p_{c-1} . Hence, all unknowns can be expressed in terms of p_0 :

$$p_r = \frac{\rho^r}{r!} p_0$$

$r = 1, 2, \dots, c$.

By insertion into (6.2) one obtains

$$p_0 = 1 / \sum_{i=0}^c \frac{\rho^i}{i!}$$

and hence,

$$p_r = \frac{\rho^r}{r!} / \sum_{i=0}^c \frac{\rho^i}{i!}$$

The probability of blocking B is the probability of finding the system in $[c]$; hence,

$$p_c = \frac{\rho^c}{c!} / \sum_{i=0}^c \frac{\rho^i}{i!} (= E_{1,c}(\rho)) \quad (6.3)$$

This is the famous *Erlang B-formula*. The function $E_{1,c}$ is called the first Erlang function.

When $c - \rho > 2\sqrt{\rho}$ the denominator in (6.3) is practically equal to e^ρ and B is nearly equal to the Poisson expression, to be denoted by ϕ :

$$B \cong \phi_c = e^{-\rho} \frac{\rho^c}{c!}.$$

